

INTRODUCTORY ECONOMETRICS

Lesson 4

Dr Javier Fernández

etpfemaj@ehu.es

Dpt. of Econometrics & Statistics

UPV—EHU

4.1 Dummy Variables. Definition and use in the GLRM.

Dummy Variables: Definition

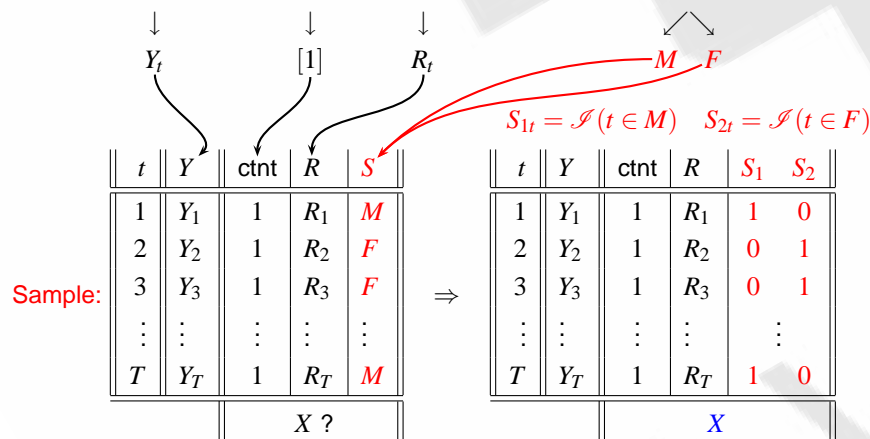
- **Qualitative explanatory var** \rightsquigarrow subsamples T_1, T_2, \dots according to **category or characteristics**
- examples:
 - ◆ **pure qualitative vars:**
 - individual diffs: sex, race, civil state, etc.
 - time diffs: season, war/peace, etc.
 - spatial diffs: countries, A.C.'s, urban/rural, etc.
 - ◆ **quantitative vars by sections:** income, age, etc.
- Recall: we cannot use qualitative vars. . . then substitute by **dummy vars.** . . .
- **Def. of Dummy Variable:**

$$D_{jt} = \begin{cases} 1, & \text{if } t \in \text{category } j; \\ 0, & \text{otherwise.} \end{cases}$$

$$\Rightarrow D_{jt} = \mathcal{I}(t \in T_j)$$

1 QV with 2 categories

Consumption = $f(\text{[cnt]}, \text{income}, \text{sex})$



In principle: substitute QV by

as many DVs as categories we have.

Dummy Var Trap: 1 qualitative var

- Model: $Y_t = \beta_0 + \beta_1 R_t + \gamma_1 S_{1t} + \gamma_2 S_{2t} + u_t$
- **Problem (Dummy Variable trap):** X is a $(T \times 4)$ matrix, but
 - $S_1 + S_2 = [1]$ (exact l.c.) $\Rightarrow \text{rk}(X) = 3 < 4$ (i.e. perfect MC)
 - $\Rightarrow \det(X'X) = 0$
 - $\Rightarrow (X'X)^{-1}$ doesn't exist!! and $\hat{\beta}$ cannot be calculated!!
- **General Solution:** eliminate **ONE** of the col's causing the problem: [1] or S_1 or S_2 .
- (POSSIBLE Solution: **eliminate intercept.** . . . but . . .

Solution: DV without a category

MOST USUAL SOLUTION: eliminate category: e.g. $F (S_2)$:

Model to estimate:

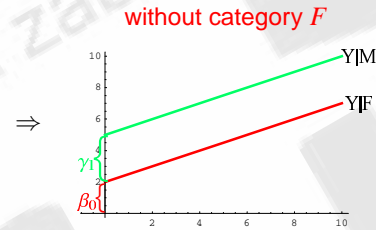
$$Y_t = \beta_0 + \beta_1 R_t + \gamma_1 S_{1t} + \gamma_2 S_{2t} + u_t$$

$$= \beta_0 + \beta_1 R_t + \gamma_1 S_{1t} + u_t$$

Subsample Models:

$$E(Y_t | S = F) = \beta_0 + \beta_1 R_t$$

$$E(Y_t | S = M) = \beta_0 + \beta_1 R_t + \gamma_1$$



$$E(Y_t | R_t = 0, S = F) = \beta_0$$

$$E(Y_t | S = M) - E(Y_t | S = F) = \gamma_1$$

Coefficient interpretation:

Coefficient Interpretation

$$E(Y_t | S = M) - E(Y_t | S = F) = \gamma_1$$

$$E(Y_t | R_t = 0, S = F) = \beta_0$$

that is,

β_0 = expected consumption Women (base) if $R_t = 0$.

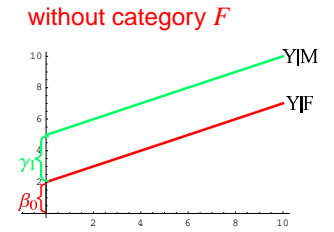
γ_1 = diff expected consumption of Men

(vs. base = Women).

β_1 = Δ consumption if $\Delta R_t = 1$ (c.p.).

Recall: This case just means different intercepts for each category.

Note: Eliminating a category \rightsquigarrow transforms it into reference base.



Usual Tests with 1 QV

Hypothesis: qualitative variable (Sex) not significant
(it doesn't affect Consumption)

i.e. M and F same Consumption:

Unrestricted Model

$$Y_t = \beta_0 + \beta_1 R_t + \gamma S_{1t} + u_t$$

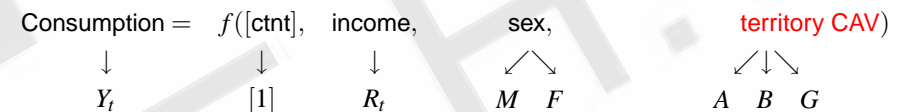
Hypothesis: $H_0 : \gamma = 0$ vs. $H_a : \gamma \neq 0$

Restricted Model:

$$Y_t = \beta_0 + \beta_1 R_t + u_t$$

Use usual t Statistic (or F Statistic based on RSS)

1 QV with 2 cats + 1 QV with 3 cats



$$S_{1t} = \mathcal{I}(t \in M)$$

$$S_{2t} = \mathcal{I}(t \in F)$$

$$T_{1t} = \mathcal{I}(t \in A)$$

$$T_{2t} = \mathcal{I}(t \in B)$$

$$T_{3t} = \mathcal{I}(t \in G)$$

Sample:

t	Y	cntn	R	S	T
1	Y_1	1	R_1	M	B
2	Y_2	1	R_2	F	G
3	Y_3	1	R_3	F	B
\vdots	\vdots	\vdots	\vdots	\vdots	\vdots
T	Y_T	1	R_T	M	A

$X ?$

t	Y	cntn	R	S_1	S_2	T_1	T_2	T_3
1	Y_1	1	R_1	1	0	0	1	0
2	Y_2	1	R_2	0	1	0	0	1
3	Y_3	1	R_3	0	1	0	1	0
\vdots	\vdots	\vdots	\vdots	\vdots	\vdots	\vdots	\vdots	\vdots
T	Y_T	1	R_T	1	0	1	0	0

X

Recall: In principle, substitute qualitative var

by as many Dummy vars as categories we have.

Dummy Var Trap: 2 qualitative vars

- Model: $Y_t = \beta_0 + \beta_1 R_t + \gamma_1 S_{1t} + \gamma_2 S_{2t} + \delta_1 T_{1t} + \delta_2 T_{2t} + \delta_3 T_{3t} + u_t$
- Problem** (DV trap):
 X is a $(T \times 7)$ matrix, but

$$S_1 + S_2 = T_1 + T_2 + T_3 = [1]$$
 (2 exact l.c.) $\Rightarrow \text{rk}(X) = 5 < 7$ (i.e. perfect MC)
- $$\Rightarrow \det(X'X) = 0$$

$$\Rightarrow (X'X)^{-1} \text{ doesn't exist!! and } \hat{\beta} \text{ cannot be calculated!!}$$
- General Solution:** eliminate **ONE** of the col's causing the problem: [1] or (S_1 or S_2) or (T_1 or T_2 or T_3).

Solution: DV without combination of categories

MOST USUAL SOLUTION:

eliminate last category of **each** DV: S_2 and T_3 :

- Model to estimate:

$$Y_t = \beta_0 + \beta_1 R_t + \gamma_1 S_{1t} + \cancel{\gamma_2 S_{2t}} + \delta_1 T_{1t} + \delta_2 T_{2t} + \cancel{\delta_3 T_{3t}} + u_t$$

$$= \beta_0 + \beta_1 R_t + \gamma_1 S_{1t} + \delta_1 T_{1t} + \delta_2 T_{2t} + u_t$$

- Subsample Models:**

	$S = M$	$S = F$	$M - F$
$T = A$	$\beta_0 + \beta_1 R_t + \gamma_1 + \delta_1$	$\beta_0 + \beta_1 R_t + \delta_1$	γ_1
$T = B$	$\beta_0 + \beta_1 R_t + \gamma_1 + \delta_2$	$\beta_0 + \beta_1 R_t + \delta_2$	γ_1
$T = G$	$\beta_0 + \beta_1 R_t + \gamma_1$	$\beta_0 + \beta_1 R_t$	γ_1
$A - G$	δ_1	δ_1	
$B - G$	δ_2	δ_2	
$A - B$	$\delta_1 - \delta_2$	$\delta_1 - \delta_2$	

Coefficient Interpretation

$$E(Y_t | S = M) - E(Y_t | S = F) = \gamma_1$$

$$E(Y_t | T = A) - E(Y_t | T = G) = \delta_1$$

$$E(Y_t | T = B) - E(Y_t | T = G) = \delta_2$$

$$E(Y_t | R_t = 0, S = F, T = G) = \beta_0$$

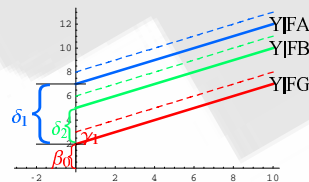
- that is,

- β_0 = expected consumption Women G (base) if $R_t = 0$.
- γ_1 = diff. expected consumption Men vs. Women .
- δ_1 = diff. expected consumption A vs. G .
- δ_2 = diff. expected consumption B vs. G .
- $\beta_1 = \Delta$ consumption if $\Delta R_t = 1$ (c.p.).

Recall: This case just means different **intercepts for each category**. Recall: Eliminating a (combination of) category(ies)

\rightsquigarrow transforms it into reference base.

without categories F nor G



Usual Tests with 2 QVs

Hypothesis: Variable Sex doesn't affect Consumption (but place of residence might do)

- Unrestricted Model:**

$$Y_t = \beta_0 + \beta_1 R_t + \gamma_1 S_{1t} + \delta_1 T_{1t} + \delta_2 T_{2t} + u_t$$

(γ_1 = diff. exp. C of M vs. F)

- Hypothesis:** $H_0 : \gamma_1 = 0$ vs. $H_a : \gamma_1 \neq 0$
- Restricted Model:**

$$Y_t = \beta_0 + \beta_1 R_t + \delta_1 T_{1t} + \delta_2 T_{2t} + u_t$$

- Use usual t Statistic (or F Statistic based on RSS)

Other usual Tests with 2 QVs

- **Unrestricted Model** (without S_2 nor T_3):

$$Y_t = \beta_0 + \beta_1 R_t + \gamma_1 S_{1t} + \delta_1 T_{1t} + \delta_2 T_{2t} + u_t$$

- ◆ Recall: γ_1 is diff. expected C of M vs. F (base)
- ◆ δ_1 and δ_2 are diff. exp. C of A and B vs. G (base)

- **Hypothesis: Same Consumption overall**

(independently of Sex and Residence):

- ◆ $H_0 : \gamma_1 = \delta_1 = \delta_2 = 0$
- ◆ **Restricted Model:**

$$Y_t = \beta_0 + \beta_1 R_t + u_t$$

- **Hypothesis: Place of Residence doesn't affect Consumption**

(but M vs. F might do):

- ◆ $H_0 : \delta_1 = \delta_2 = 0$
- ◆ **Restricted Model:**

$$Y_t = \beta_0 + \beta_1 R_t + \gamma_1 S_{1t} + u_t$$

Other usual Tests with 2 QVs

- **Unrestricted Model** (without S_2 nor T_3):

$$Y_t = \beta_0 + \beta_1 R_t + \gamma_1 S_{1t} + \delta_1 T_{1t} + \delta_2 T_{2t} + u_t$$

- ◆ Recall: δ_1 and δ_2 are diff. expected C of A and B vs. G

(base)

- **Hypothesis: Residents of same sex in A and B have same consumption level** (but G might be different):

- ◆ $H_0 : \delta_1 = \delta_2$ vs. $H_a : \delta_1 \neq \delta_2$
- ◆ **Restricted Model:**

$$Y_t = \beta_0 + \beta_1 R_t + \gamma_1 S_{1t} + \delta \underbrace{(T_{1t} + T_{2t})}_{1-T_{3t}} + u_t$$

- **Hypothesis: Residents of same sex in B and G have same consumption level** (but A might be different):

- ◆ $H_0 : \delta_2 = 0$ vs. $H_a : \delta_2 \neq 0$
- ◆ **Restricted Model:**

$$Y_t = \beta_0 + \beta_1 R_t + \gamma_1 S_{1t} + \delta_1 T_{1t} + u_t$$

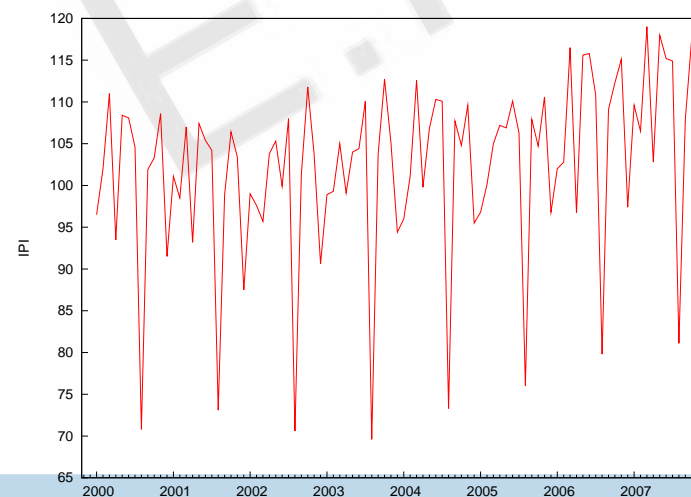
4.2 Seasonal effects

Seasonal effect

- **Seasonal effect:**

- **Seasonal var** \rightsquigarrow subsamples T_1, T_2, \dots

according to **seasons/months**



Seasonal Dummy Variables: Definition

- Def. of Seasonal Dummy Variable:

$$D_{jt} = \begin{cases} 1, & \text{if } t \in \text{season } j = 1, 2, 3, 4, \dots; \\ 0, & \text{otherwise.} \end{cases}$$

date (t)	IPI_t	X_t	D_{1t}	D_{2t}	D_{3t}	D_{4t}
1975.1	-	-	1	0	0	0
1975.2	-	-	0	1	0	0
1975.3	-	-	0	0	1	0
1975.4	-	-	0	0	0	1
1976.1	-	-	1	0	0	0
1976.2	-	-	0	1	0	0
1976.3	-	-	0	0	1	0
1976.4	-	-	0	0	0	1
1977.1	-	-	1	0	0	0
...
2000.1	-	-	1	0	0	0
2000.2	-	-	0	1	0	0
2000.3	-	-	0	0	1	0
2000.4	-	-	0	0	0	1
2001.1	-	-	1	0	0	0
...

- e.g. for quarterly data:

Seasonal Dummy Variables: Definition (2)

- Model to estimate:

$$IPI_t = \beta_0 + \beta_1 X_t + \gamma_1 D_{1t} + \gamma_2 D_{2t} + \gamma_3 D_{3t} + \gamma_4 D_{4t} + u_t$$

$$= \beta_0 + \beta_1 X_t + \gamma_1 D_{1t} + \gamma_2 D_{2t} + \gamma_3 D_{3t} + u_t$$

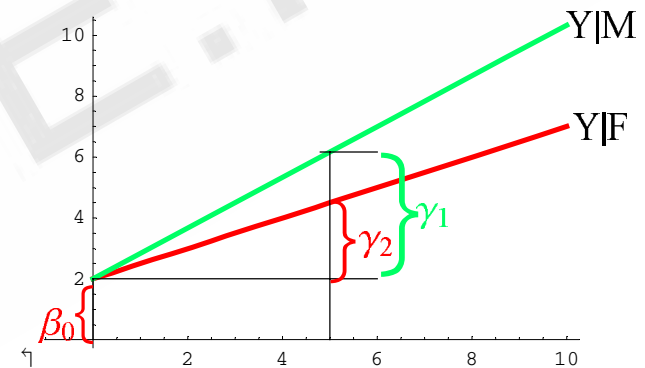
- interpretation of γ parameters?
- What if data are monthly observations (as in the IPI example actually)?

date (t)	IPI_t	X_t	D_{1t}	D_{2t}	D_{3t}	D_{4t}	$D_{1t} \dots$	$\dots D_{12t}$
1975.jan	-	-	1	0	0	0	1 0 0 0	0 0 0 0
1975.feb	-	-	1	0	0	0	0 1 0 0	0 0 0 0
1975.mar	-	-	1	0	0	0	0 0 1 0	0 0 0 0
1975.apr	-	-	0	1	0	0	0 0 0 1	0 0 0 0
1975.may	-	-	0	1	0	0	0 0 0 0	1 0 0 0
1975.jun	-	-	0	1	0	0	0 0 0 0	0 0 1 0
1975.jul	-	-	0	0	1	0	0 0 0 0	0 0 0 1
1975.ago	-	-	0	0	1	0	0 0 0 0	0 0 0 0
1975.sep	-	-	0	0	1	0	0 0 0 0	0 0 0 0
1975.oct	-	-	0	0	0	1	0 0 0 0	0 0 0 0
1975.nov	-	-	0	0	0	1	0 0 0 0	0 0 0 0
1975.dec	-	-	0	0	0	1	0 0 0 0	0 0 0 0
1976.jan	-	-	1	0	0	0	1 0 0 0	0 0 0 0
1976.feb	-	-	1	0	0	0	0 1 0 0	0 0 0 0
1976.mar	-	-	1	0	0	0	0 0 1 0	0 0 0 0
...

4.3 Interaction between DVs and quantitative Vars

Interaction between DVs and quantitative Vars

Instead of different *intercepts*, we require **different slopes** for each category:



that is, different **response** "Y" for same "X"

Dummy Var Trap: interaction

- Matrix X :

cnt	R	$R \times S_1$	$R \times S_2$
1	R_1	$R_1 \times 1$	$R_1 \times 0$
1	R_2	$R_2 \times 0$	$R_2 \times 1$
1	R_3	$R_3 \times 0$	$R_3 \times 1$
\vdots	\vdots	\vdots	\vdots
1	R_T	$R_T \times 1$	$R_T \times 0$

 \Rightarrow

cnt	R	RS_1	RS_2
1	R_1	R_1	0
1	R_2	0	R_2
1	R_3	0	R_3
\vdots	\vdots	\vdots	\vdots
1	R_T	R_T	0

- Model:

$$Y_t = \beta_0 + \beta_1 R_t + \gamma_1 R_t S_{1t} + \gamma_2 R_t S_{2t} + u_t$$

- Problem** (DV trap): X is $T \times 4$, but

$$RS_1 + RS_2 = R \Rightarrow \text{rk}(X) = 3 < 4 \quad (\text{exact MultiCol!})$$

- General Solution:** eliminate **ONE** of the col's causing the problem: R or RS_1 or RS_2 .

Solution: Interaction without a category

- MOST USUAL SOLUTION:**

eliminate last category of the DV: F (RS_2):

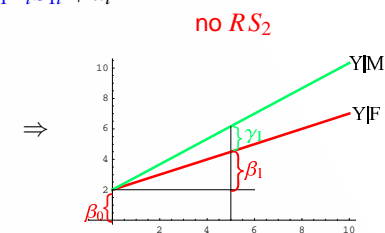
- Model to estimate:

$$Y_t = \beta_0 + \beta_1 R_t + \gamma_1 R_t S_{1t} + u_t$$

- Subsample Models:**

$$E(Y_t | S = F) = \beta_0 + \beta_1 R_t$$

$$E(Y_t | S = M) = \beta_0 + \underbrace{(\beta_1 + \gamma_1)}_{\beta_1^*} R_t$$



$$E(Y_t | R_t = 0) = \beta_0$$

$$\frac{\partial E(Y_t | S = F)}{\partial R_t} = \beta_1$$

$$\frac{\partial E(Y_t | S = M)}{\partial R_t} = \beta_1 + \gamma_1$$

- Coefficient interpretation:**

Coefficient Interpretation

$$E(Y_t | R_t = 0) = \beta_0$$

$$\frac{\partial E(Y_t | S = F)}{\partial R_t} = \beta_1$$

$$\frac{\partial E(Y_t | S = M)}{\partial R_t} = \beta_1 + \gamma_1$$

- that is,

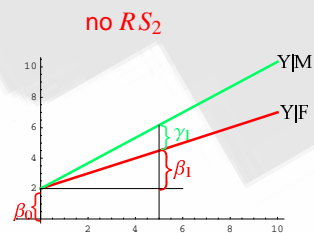
β_0 = expected consumption if $R_t = 0$.

β_1 = Δ consumption Women if $\Delta R_t = 1$ (c.p.).

γ_1 = diff Δ consumption for Men (vs. base = Female).

Recall: This case means **different slopes for each category**. Recall: again eliminating a category \rightsquigarrow

transforms it into reference base.



Usual Tests with Interaction

Hypothesis: M and F equal Consumption or variable Sex doesn't affect Consumption:

- Unrestricted Model:**

$$Y_t = \beta_0 + \beta_1 R_t + \gamma_1 R_t S_{1t} + u_t$$

- Hypothesis:** $H_0 : \gamma_1 = 0$ vs. $H_a : \gamma_1 \neq 0$

- Restricted Model:**

$$Y_t = \beta_0 + \beta_1 R_t + u_t$$

- Use usual t Statistic (or F Statistic based on RSS)

THE END

eman ta zabal zazu

IP.V. E.H.U.